Adaptive matrix distances aiming at optimum regression subspaces

M. Strickert[a*], Axel J. Soto[bc], and Gustavo E. Vazquez[b]

[a] Institute for Vision and Graphics (IVG), University of Siegen, Germany
[b] Laboratory for Research and Development in Scientific Computing, DCIC,
Universidad Nacional del Sur Bahía Blanca, Argentina
[c] Planta Piloto de Ingeniería Química, CONICET-UNS, Argentina
[*] Corresponding author: `marc.strickert@uni-siegen.de`

Abstract.   A new supervised adaptive metric approach is introduced for mapping an input vector space to a plottable low-dimensional subspace in which the pairwise distances are in maximum correlation with distances of the associated target space. The formalism of multivariate subspace regression (MSR) is based on cost function optimization, and it allows assessing the relevance of input vector attributes. An application to molecular descriptors in a chemical compound data base is presented for targeting octanol-water partitioning properties.

Keywords. Data-driven metric, feature rating, informative subspace.

1   Introduction

The connection of data vectors with a specific target is a fundamental problem in data analysis. Input data of real valued vectors are basic entities in many scientific fields, for example, ranging from spectrum data and gene expression data in medicine and biology via sensor measurements in engineering sciences to compound fingerprints in chemistry. The target can be a categorial label in classification tasks, a real-valued dependent variable in regression problems or even a vector of properties in association scenarios.

The empirical assessment of target information is often a time consuming and expensive task, for example, the identification of tissue types in histological samples requires manual work and wet-lab experiments. Due to this careful intervention it can be assumed that the targets assigned to the sample vectors reflect a reliable and immutable ground truth, up to a few mislabelings. In contrast to this, the data vectors live in a space of measurements that usually quantify general properties, but which should, preferably, be predictive of the targets.

A number of different techniques exists that allow a link between the input space and the target space, such as linear discriminant analysis (LDA) for discrete class labels [3], generalized linear models (GLM) for regression tasks [2], and canonical correlation analysis (CCA) for association problems [1]. These are well-established linear models. Complementary, neural networks like feedforward networks provide a nonlinear connection between input space and the target, but they do require a choice of architectural parameters in the hidden

layer or the selection of an appropriate learning algorithm, making it difficult to assess stability and reliability.

The approach presented here allows application to discrete labels and multivariate regression variables. It follows the assumption that the input space is adapted in such a way to optimize the input vector representations for matching the target relationships. More precisely, here a matrix distance with the structure of the Mahalanobis distance is adapted to yield maximum correlation of the pairwise vector distances and the associated pairwise target distances. For this specific distance, the approach can be interpreted as an alternative way for solving linear inverse models, such as calculated by the Moore-Penrose pseudoinverse. The model will be called multivariate subspace regression (MSR) in the following.

Adaptive matrix metrics have been proposed to be useful in k-nearest neighbors [8] and learning vector quantization with local metrics [4]. Recently a feature ranking method based on a class discriminant function has been proposed as robust alternative to LDA [6]. It has been shown this method is useful for complementing hard feature selection strategies of evolutionary algorithms (EA) for assessing molecular descriptors for biological and physicochemical property prediction essential in drug design [5, 7]. Yet, this metric-driven feature rating scheme required a simplifying two-class assumption of low and high octanol-water partitioning coefficient (logP) targets. Since the new formalism presented in the following allows to deal with feature selection in general regression contexts, the data base of 439 chemical compounds is revisited to study the influence of the underlying 73 molecular descriptors on the logP regression task.

2 Methods

Let N input vectors be given as $\mathbf{x}^j \in \mathbf{X} \subset \mathbb{R}^M, \mathbf{x}^j = (x_k^j)_{k=1...M}, 1 \leq j \leq N$ with associated target vectors $\boldsymbol{l}^j \in \mathbf{L} \subset \mathbb{R}^q, \boldsymbol{l}^j = (l_k^j)_{k=1...q}$. The transformable input space $\mathbf{X}$ shall be linked to the immutable target space $\mathbf{L}$ by the relationship

$$S_{\mathrm{d}^v} = \mathsf{r}(\mathbf{D_L}, \mathbf{D_X^{\boldsymbol{\lambda}}}) = \max. \tag{1}$$

Therein, $\mathbf{D_L}$ is the distance matrix of all pairs of target vectors, here defined by Euclidean distance; $\mathbf{D_X^{\boldsymbol{\lambda}}}$ is the matrix of all input vector distances which do depend on the parameter vector $\boldsymbol{\lambda}$. Thus, parameters are sought that maximize the Pearson correlation (r) between input and target space.

The model parameters are obtained by maximizing the functional S_{d^v} using its gradient

$$\frac{\partial S_{\mathrm{d}^v}}{\partial \boldsymbol{\lambda}} = \frac{\partial \mathsf{r}(\mathbf{D_L}, \mathbf{D_X^{\boldsymbol{\lambda}}})}{\partial \mathbf{D_X^{\boldsymbol{\lambda}}}} \cdot \frac{\partial \mathbf{D_X^{\boldsymbol{\lambda}}}}{\partial \boldsymbol{\lambda}} = \sum_{i=1}^{N} \sum_{j=1}^{N} \frac{\partial \mathsf{r}(\mathbf{D_L}, \mathbf{D_X^{\boldsymbol{\lambda}}})}{\partial (\mathbf{D_X^{\boldsymbol{\lambda}}})_{i,j}} \cdot \frac{\partial (\mathbf{D_X^{\boldsymbol{\lambda}}})_{i,j}}{\partial \boldsymbol{\lambda}}. \tag{2}$$

The required derivatives of the Pearson correlation are calculated by:

$$\frac{\partial \mathsf{r}(\mathbf{D_L}, \mathbf{D_X^{\boldsymbol{\lambda}}})}{\partial (\mathbf{D_X^{\boldsymbol{\lambda}}})_{i,j}} = \frac{\left((\mathbf{D_L})_{i,j} - \mu_{\mathbf{D_L}}\right) - \frac{\mathscr{B}}{\mathscr{D}} \cdot \left((\mathbf{D_X^{\boldsymbol{\lambda}}})_{i,j} - \mu_{\mathbf{D_X^{\boldsymbol{\lambda}}}}\right)}{\sqrt{\mathscr{C} \cdot \mathscr{D}}} . \tag{3}$$

Therein, $\mu_{\mathbf{D_L}}$ and $\mu_{\mathbf{D_X^\lambda}}$ denote the mean values of the matrices, and the notations $\mathscr{B} = \sum_{i=1}^{N}\sum_{j=1}^{N}\big((\mathbf{D_L})_{i,j}-\mu_{\mathbf{D_L}}\big)\cdot\big((\mathbf{D_X^\lambda})_{i,j}-\mu_{\mathbf{D_X^\lambda}}\big)$, $\mathscr{C} = \sum_{i=1}^{N}\sum_{j=1}^{N}\big((\mathbf{D_L})_{i,j}-\mu_{\mathbf{D_L}}\big)^2$ and $\mathscr{D} = \sum_{i=1}^{N}\sum_{j=1}^{N}\big((\mathbf{D_X^\lambda})_{i,j}-\mu_{\mathbf{D_X^\lambda}}\big)^2$ are used.

For optimization the quasi Newton Broyden-Fletcher-Goldfarb-Shanno method was taken. Optimization is stopped, if the improvement of subsequent evaluations of S_{d^v} drops below 10^{-8}.

Because of its flexibility, the input vectors $\boldsymbol{x}^i$ and $\boldsymbol{x}^j \in \mathbf{X}$ are chosen to be compared by a matrix metric with Mahalanobis structure in this work:

$$(\mathbf{D_X^\lambda})_{i,j} = \mathrm{d}^v\left(\boldsymbol{x}^i, \boldsymbol{x}^j|\boldsymbol{\lambda}\right) = \sqrt{(\boldsymbol{x}^i - \boldsymbol{x}^j)^\top \cdot \boldsymbol{\lambda} \cdot \boldsymbol{\lambda}^\top \cdot (\boldsymbol{x}^i - \boldsymbol{x}^j)}. \tag{4}$$

Unlike Mahalanobis distance there is no inverse covariance matrix employed. Instead, the outer self-product of the parameter matrix $\boldsymbol{\lambda} \in \mathbb{R}^{M\times u}$ defines an adaptive matrix $\boldsymbol{\Lambda} = \boldsymbol{\lambda}\cdot\boldsymbol{\lambda}^\top$. This positive-definite matrix $\boldsymbol{\Lambda}$ contains components that weigh the influence of attribute pairs (g, k) in the data space. Its maximum rank is u if the number of input dimensions M is larger than the u-dimensional subspace defined by $\mathbf{X}^\top\cdot\boldsymbol{\lambda}$. This subspace is an informative representation of the input space focused on the target association. Since, in principle, any dimension u can be chosen it is more flexible than inverse linear models which require the same dimensionality as the target space. As a very general recommendation, a choice of $u \leq M$ and $u \leq N$, or $u \leq 3$ for visualization is possible, depending on the desired representation accuracy expressed by S_{d^v}.

The derivative of Eqn. 4, useful for optimization, is

$$\frac{\partial \mathrm{d}^v\left(\boldsymbol{x}^i, \boldsymbol{x}^j|\boldsymbol{\lambda}\right)}{\partial \boldsymbol{\lambda}} = \frac{(\boldsymbol{x}^i - \boldsymbol{x}^j) \cdot \left((\boldsymbol{x}^i - \boldsymbol{x}^j)^\top \cdot \boldsymbol{\lambda}\right)}{\mathrm{d}^v\left(\boldsymbol{x}^i, \boldsymbol{x}^j|\boldsymbol{\lambda}\right)}\,. \tag{5}$$

If regression targets are modeled in a one-dimensional subspace $\boldsymbol{p} = \mathbf{X}^\top\cdot\boldsymbol{\lambda}$, the projected scalar values obviously depend on the data vectors and the parameter vector. Arbitrary scaling and shifting of the projections $\boldsymbol{p}$ are matched to fit by choosing α and β in $\hat{\boldsymbol{p}} = \alpha \cdot \boldsymbol{p} + \beta$ such that

$$F = \sum_{i=1}^{N}\left(l_i - (\alpha \cdot p_i + \beta)\right)^2 = \min. \tag{6}$$

3   Results

A compound data set with 73 molecular features and associated logP values for 439 chemical compounds has been taken for the analysis, online available at `http://dig.ipk-gatersleben.de/sardux/sardux.html` [7]. Therein, an independent test set of 30 compounds has been defined that covers the range of logP values uniformly and that is not confined in the convex hull of the training data.

Two relevant cases are considered here: a multidimensional regression task on the scalar logP target values and a regression involving three disjoint classes.

While the first application shall illustrate its competitiveness with state-of-the-art inverse linear models, the second application unfolds its unique use for mapping data related to the three-dimensional space of independent (orthogonal) class labels onto a two-dimensional subspace.

The only model parameter needed to be chosen is the dimensionality of the subspace, i.e. one and two in these examples. The stability is assessed by running the optimization 10 times initialized independently with random parameter vectors $\boldsymbol{\lambda}$.

For the multidimensional regression problem, well-proven tools are available for comparison to the averages and standard deviations of 10 MSR runs: the matrix left division operator '\' based on MATLAB Householder reflections and the R:limSolve package implementing the Moore-Penrose pseudoinverse. The

r^2	MSR	MATLAB(7.5.0):'\'	R:limSolve(11.09)
train	0.9357 ± 0.0001	0.9231	0.9361
test	0.8704 ± 0.0004	0.8413	0.8660

Table 1: Regression results of the new method compared to approaches based on pseudoinverse calculations.

comparison Table 1 shows that MSR is better than MATLAB, and slightly worse then R:limSolve only for the training data. The low standard deviations of MSR indicate a very good reproducibility. The left panel of Figure 1 shows the MSR regression result of a model of median performance on the training data, using projections transformed according to Eqn. 6.

The three disjoint class problem has been created by splitting the logP values into the lower, middle, and upper 33.3% quantile, assigning three-dimensional targets $(0, 0, 1)$ for logP < 1.78, $(0, 1, 0)$ for $1.78 \leq$ logP < 3.0132, and $(1, 0, 0)$ for logP ≥ 3.0132 values. Note that this is different from assigning integer class labels 1, 2, and 3, which, for example, would induce a closer relationship of the class labels 1 and 2 rather than 1 and 3. The right panel of Figure 1 shows the two-dimensional transformation of the data space aiming at arranging the projections according to the target relationships. Despite of logP being a continuous regression variable naturally reflected in the molecular descriptor vectors, by exploiting the 73 dimensions of the input vectors, MSR is able to provide a good separation of the projections with only decent overlap.

Figure 2 shows the attribute relevance profiles corresponding to the two regression tasks. At first glance a high degree of similarity can be detected, such as the highly important molecular van der Waals volume (Mv). Yet, descriptors like atomic polarizability (Sp) and the number of sulfor atoms(nS) show quite a different influence on the specific task. These results are quite certain, because the box plots display a high reproducibility of the model runs. As interesting to chemists, the profiles do also indicate that most variables do only have minor relevance on the regression tasks.

logP regression
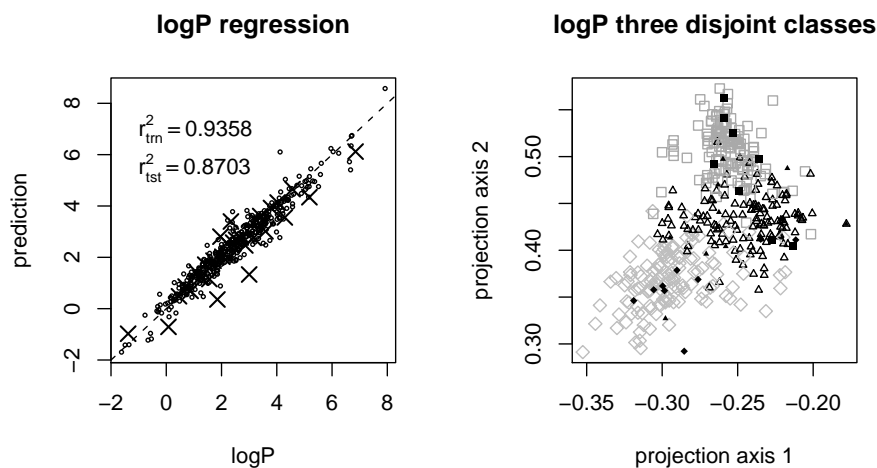

logP three disjoint classes

Fig. 1: MSR projection subspaces of logP data set. Left: 1D regression task (big crosses: test data). Right: 3D disjoint label separation in 2D (filled symbols: test data).
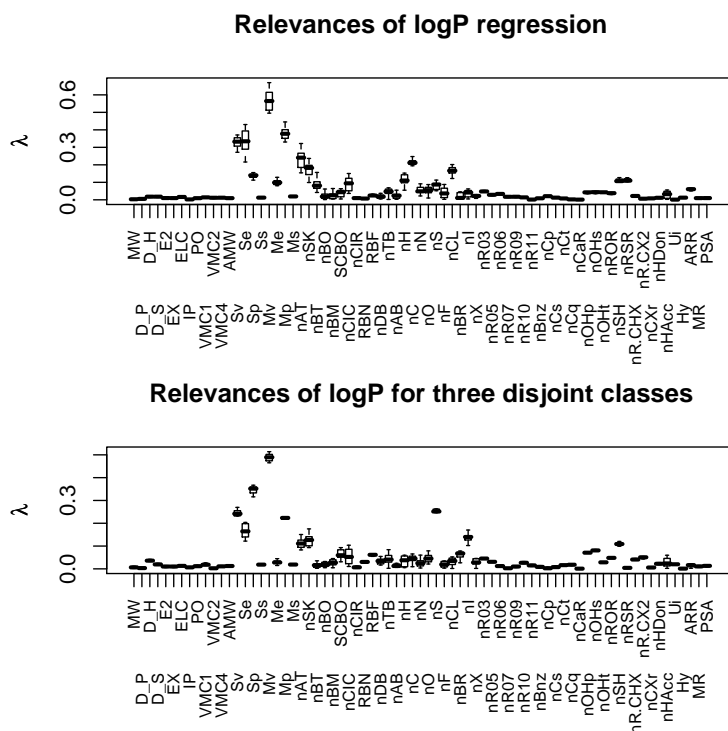
Relevances of logP regression

Relevances of logP for three disjoint classes

Fig. 2: Molecular descriptor relevance profiles given by $\sum_{i=1}^{u} |\boldsymbol{\lambda}^i - \mu_{\lambda^i}|$. Top: 1D regression task ($u = 1$). Bottom: 3D disjoint label separation (u=2).

4 Conclusions

The proposed method adjusts a data metric of Mahalanobis structure for arranging the input vector relationships in good agreement to the target relationships. The metric parameters result from the optimization of a correlation-based cost function connecting input and target space. The distance can be re-interpreted as a mapping of the data vectors to a low-dimensional Euclidean space where points aim at reflecting the target relationships.

These transformed data points can be used as data replacement in subsequent analysis steps with standard Euclidean methods for classification and multivariate regression. In contrast to traditional feature assessment methods, the proposed adaptive matrix metric contains information not only about singular attributes, but about pairs of attributes. This is, for example, useful in combination with feedforward neural networks, because they integrate over input feature combinations in the hidden layer.

Alternatively, the learned metric parameters can be used for identifying the relevance of pairs of input data attributes. As demonstrated for the logP prediction task, the rating may depend on the target of regression or multiple class labeling. The method has good empirical convergence properties and good potential for general data processing tasks.

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